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A control policy for a priority queue with removable server ^{*})

H.C. Tijms

ABSTRACT

This paper studies a control policy for an M/G/1 non-preemptive priority queueing system with removable server and two priority classes. This policy turns off the server when the system is empty and turns on the server when a given linear combination of the numbers of class 1 and class 2 customers in the system exceeds a certain value. Expressions for the long-run average numbers of class 1 and class 2 customers in the system are derived.

^{*}) This paper is not for review; it is meant for publication in a journal.

Consider a service station with a single server at which customers of classes 1 and 2 arrive in accordance with independent Poisson processes with rates λ_1 and λ_2 , respectively. Customers of class 1 have non-preemptive priority over customers of class 2. The order in which customers of a given priority class are served is immaterial in our considerations assuming that this order is independent of the service times. A customer of class i will be called an i -customer, $i=1,2$. Let the service times of different customers be independent random variables with finite first moment μ_i and finite second moment $\mu_i^{(2)}$ for i -customers. Let $\rho_i = \lambda_i \mu_i$. It is assumed that $\rho < 1$, where $\rho = \rho_1 + \rho_2$. The policy for controlling the system is to turn off the server only when the system is empty and to turn on the server at the first time when the sum of α_1 times the number of 1-customers present and α_2 times the number of 2-customers present exceeds β , where α_1 , α_2 and β are non-negative constants with $\alpha_1 + \alpha_2 > 0$. We call this policy an $(\alpha_1, \alpha_2, \beta)$ -policy.

For a single class of customers, such a policy was studied by BALACHANDRAN[1], BELL[2], HEYMAN[4], YADIN AND NAOR[10], and others. The $(\alpha_1, \alpha_2, \beta)$ -policy for the priority model was studied by Bell[3]. Assuming that the service time distributions do not differ for the two classes of customers and that there is a linear cost structure, he proved that an average cost optimal policy exists which is of the $(\alpha_1, \alpha_2, \beta)$ -type.

The purpose of this paper is to derive an expression for the average number of i -customers in the system, $i=1,2$. Our derivation will be based on the theory of regenerative processes, cf. STIDHAM[7]. Finally, we superimpose a linear cost structure and determine the best $(1,1,\beta)$, $(1,0,\beta)$ and $(0,1,\beta)$ -policies with respect to the average cost criterion.

The results in this paper can be extended without difficulty to cover set-up times and close-down times as in Yadin and Naor[10].

1. NOTATION AND PRELIMINARY RESULTS

For convenience we assume throughout this paper that the server is turned off at epoch 0, so no customers are present at epoch 0. Let X be the next epoch at which the server is turned off. For any $t \geq 0$ and $i=1,2$, let $L_i(t)$ be the number of i -customers in the system at time t (including the i -customer being served at time t , if any). Observe that $\int_0^t L_i(s)ds$ represents the total time spent by i -customers in the system during $(0,t]$. By the memoryless property of the Poisson process, any epoch at which the server is turned off is a regeneration epoch for the process $\{L_i(t), t \geq 0\}, i=1,2$. Let a *cycle* be the time interval between two successive epochs at which the server is turned off. We shall see in section 3 that both X and $\int_0^X L_i(s)ds$ have a finite expectation. Hence, by the theory of regenerative processes (p.99 in ROSS[6] and Theorem 1 of Stidham[7]),

$$L^{(i)} = \lim_{t \rightarrow \infty} (1/t) E\left\{\int_0^t L_i(s)ds\right\}$$

exists and is finite for $i=1,2$. Also, the random variable $(1/t)\int_0^t L_i(s)ds$ converges with probability 1 to $L^{(i)}$ as $t \rightarrow \infty$. Furthermore, for $i=1,2$,

$$L^{(i)} = E\left\{\int_0^X L_i(s)ds\right\}/EX, \quad (1)$$

i.e., the long-run average number of i -customers in the system is equal to the quotient of the expected total time spent by i -customers in the system during one cycle and the expected length of one cycle.

Remark. Let the average wait of an i -customer in the system be defined by $W^{(i)} = \lim_{n \rightarrow \infty} (1/n) E\left\{\sum_{k=1}^n W_{ki}\right\}$, where W_{ki} denotes the time spent by the k th i -customer in the system (including his service time). Since the expected number of i -customers served during one cycle is finite, we have by Theorem 2 of Stidham[7] that $W^{(i)}$ is well defined and finite for $i=1,2$. Similarly,

we may define $L_q^{(i)}$ and $W_q^{(i)}$, where $L_q^{(i)}$ represents the average number of i -customers in the queue and $W_q^{(i)}$ represents the average wait of an i -customer in the queue. We have $L_q^{(i)} = \lambda_i W_q^{(i)}$ and $L_q^{(i)} = \lambda_i W_q^{(i)}$ for $i=1,2$ (see JEWELL[5] and Stidham[8]). Since $W_q^{(i)} = W_q^{(i)} - \mu_i$, we have $L_q^{(i)} = L_q^{(i)} - \rho_i$, $i=1,2$.

Let $\lambda = \lambda_1 + \lambda_2$. Observe that if we lump the two separate arrival processes together, the superimposed process is a Poisson process with rate λ . For $i=1,2$, let $p_i = \lambda_i / \lambda$, so that p_i is the probability that an arbitrary customer is an i -customer. Let the probability distribution function $F(x)$ be defined as follows: in case $\alpha_1 \neq \alpha_2$ the points α_1 and α_2 are points of increase of F with weights p_1 and p_2 ; otherwise, the point $\alpha_1 (= \alpha_2)$ is a point of increase of F with weight 1. Denote by $F^n(x)$ the n -fold convolution of F with itself, and let $M(x) = \sum_{n=1}^{\infty} F^n(x)$, $x \geq 0$. The renewal function $M(x)$ is the unique solution which is bounded on finite intervals to

$$M(x) = F(x) + \int_0^x F(x-y) dM(y), \quad x \geq 0, \quad (2)$$

(p.35 in Ross[6]). For any $t \geq 0$, let $N_i(t)$ be the number of i -customers arriving in $(0, t]$. For any $x \geq 0$, let

$$T(x) = \inf\{t \mid \sum_{i=1}^2 \alpha_i N_i(t) > x\}, \quad v_i(x) = N_i(T(x)), \quad W_i(x) = \int_0^{T(x)} L_i(s) ds \quad (i=1,2).$$

Given that an (α_1, α_2, x) -policy is used, $v_i(x)$ represents the number of i -customers in the system at the first epoch at which the server is turned on, and $W_i(x)$ represents the total time spent by i -customers in the system up to that epoch. Let $v(x) = v_1(x) + v_2(x)$, $x \geq 0$. For any $x \geq 0$, let

$$\begin{aligned} a_i(x) &= E v_i(x), \quad b_i(x) = E\{v_i(x)(v_i(x)-1)\}, \quad w_i(x) = E W_i(x) \quad (i=1,2), \\ a(x) &= E v(x), \quad b(x) = E\{v(x)(v(x)-1)\}, \quad c(x) = E\{v_1(x)v_2(x)\}. \end{aligned}$$

THEOREM 1. For $i=1,2$, let $a_i(u) = 0$ for $u < 0$, and let $g_i(x) = p_i + 2p_i a_i(x - \alpha_i)$ for $x \geq 0$. Then, for any $x \geq 0$,

$$a_i(x) = p_i \{1 + M(x)\} \quad \text{for } i=1,2, \quad (3)$$

$$b_i(x) = g_i(x) + \int_0^x g_i(x-y) dM(y) - a_i(x) \quad \text{for } i=1,2, \quad (4)$$

$$a(x) = 1 + M(x), \quad b(x) = 2M(x) + 2 \int_0^x M(x-y) dM(y), \quad (5)$$

$$c(x) = (1/2) \{b(x) - b_1(x) - b_2(x)\}. \quad (6)$$

Let $a(u)=0$ for $u<0$, and let $h_i(x)=(1/\lambda)p_i a(x-\alpha_i)$ for $x \geq 0$ and $i=1,2$. Then,

$$w_i(x) = h_i(x) + \int_0^x h_i(x-y) dM(y) \quad \text{for } x \geq 0 \text{ and } i=1,2. \quad (7)$$

Proof. Let T be the arrival epoch of the first customer and fix i . We can write $v_i(x) = M + N$, where $M=1$ if the first customer is an i -customer and $M=0$ otherwise, and N denotes the number of i -customers arriving in $(T, T(x)]$.

Clearly, under the condition that the first customer is an j -customer, the random variable N has the same distribution as $v_i(x-\alpha_j)$, where $v_i(u)=0$ for $u<0$. Now, $a_i(x) = p_i + p_1 a_i(x-\alpha_1) + p_2 a_i(x-\alpha_2)$, so

$$a_i(x) = p_i + \int_0^x a_i(x-y) dF(y), \quad x \geq 0.$$

This is a renewal equation whose unique solution is given by (3). Let

$d_i(x) = E[v_i(x)]^2$. Using $[v_i(x)]^2 = M^2 + 2MN + N^2$, we obtain

$$d_i(x) = g_i(x) + \int_0^x d_i(x-y) dF(y), \quad x \geq 0.$$

The unique solution of this renewal equation is given by the sum of the first two terms from the right side of (4). In the same way the relation (5) can be derived, while (6) follows from $[v(x)]^2 = [v_1(x)]^2 + 2v_1(x)v_2(x) + [v_2(x)]^2$.

By $ET(y) = (1/\lambda)Ev(y)$, we have $ET(y) = (1/\lambda)a(y)$ for $y \geq 0$. Using this and considering the waiting time of the first customer and that of the next customers separately, it follows as above that $w_i(x)$ satisfies the renewal equation $w_i(x) = h_i(x) + \int_0^x w_i(x-y) dF(y)$ for $x \geq 0$, so $w_i(x)$ is given by (7). This ends the proof.

We note that, by (2) and (5), $h_1(x) + h_2(x) = M(x)/\lambda$ for $x \geq 0$.

2. BASIC RESULT

In this section we shall give a result which will be basic in our derivation of an expression for $L^{(i)}$, $i=1,2$.

Denote by $S(n_1, n_2)$ the time elapsed from the start of a service when n_1 1-customers and n_2 2-customers are in the system until the next epoch at which the system is empty. Let $s(n_1, n_2) = ES(n_1, n_2)$, and let $u_i(n_1, n_2)$ be the expected total time spent by i -customers in the system during the time $S(n_1, n_2)$, $i=1,2$ and $n_1, n_2=0,1,\dots$.

It is routine to prove the next theorem (see pp.6-7, 9-11 in Tijms[9]).

THEOREM 2. For $n_1, n_2=0,1,\dots$ with $n_1+n_2>0$,

$$\begin{aligned} s(n_1, n_2) &= t_b n_1 + \{n_2 + \lambda_2 t_b n_1\} t_{b2}, \\ u_1(n_1, n_2) &= w_1 n_1 + (1/2) t_b n_1 (n_1 - 1) + \{\lambda_2 t_b n_1 + n_2\} u_1(0, 1), \\ u_2(n_1, n_2) &= t_b (1 + \lambda_2 t_{b2}) n_1 n_2 + \{\lambda_2 t_b n_1 + n_2\} u_2(0, 1) + \\ &\quad + (1/2) \{\lambda_2 + \lambda_2^2 t_{b2}\} \{t_b^{(2)} n_1 + t_b^2 n_1 (n_1 - 1)\} + (1/2) t_{b2} n_2 (n_2 - 1), \end{aligned}$$

where

$$\begin{aligned} t_b &= \mu_1 / (1 - \rho_1), \quad t_b^{(2)} = \mu_1^{(2)} / (1 - \rho_1)^3, \quad t_{b2} = \mu_2 / (1 - \rho), \\ w_1 &= \mu_1 / (1 - \rho_1) + \lambda_1 \mu_1^{(2)} / 2(1 - \rho_1)^2, \\ u_1(0, 1) &= (1 - \rho)^{-1} [\mu_1^{(2)} \lambda_1 \mu_2 / 2(1 - \rho_1) + \mu_2^{(2)} \lambda_1 / 2 + \rho_1 \mu_2], \\ u_2(0, 1) &= (1 - \rho)^{-1} [\mu_1^{(2)} \lambda_1 \rho_2 / 2(1 - \rho_1)(1 - \rho) + \mu_2^{(2)} \lambda_2 / 2(1 - \rho) + (1 - \rho_1) \mu_2]. \end{aligned}$$

3. THE AVERAGE NUMBER OF i -CUSTOMERS IN THE SYSTEM

THEOREM 3. For $i=1,2$, let $C_i = \rho_i + \lambda_i [2(1 - \rho_1)(1 - \delta_i)]^{-1} [\lambda_1 \mu_1^{(2)} + \lambda_2 \mu_2^{(2)}]$, where $\delta_1=0$ and $\delta_2=\rho$. Then, for any $(\alpha_1, \alpha_2, \beta)$ -policy,

$$L^{(1)} = C_1 + \lambda(1 - \rho) [1 + M(\beta)]^{-1} [w_1(\beta) + \mu_1 b_1(\beta) / 2(1 - \rho_1)], \quad (8)$$

$$\begin{aligned} L^{(2)} &= C_2 + \lambda(1 - \rho) [1 + M(\beta)]^{-1} [w_2(\beta) + \mu_1 c(\beta) / (1 - \rho) + \mu_2 b_2(\beta) / 2(1 - \rho) + \\ &\quad + \lambda_2 \mu_1^2 b_1(\beta) / 2(1 - \rho_1)(1 - \rho)]. \end{aligned} \quad (9)$$

Proof. Using the Theorems 1 and 2, it is readily verified that the expected length of one cycle equals

$$(1/\lambda)[1+M(\beta)]+Es(v_1(\beta),v_2(\beta)) = [\lambda(1-\rho)]^{-1}[1+M(\beta)].$$

The expected total time spent by i -customers in the system during one cycle equals $w_i(\beta)+Eu_i(v_1(\beta),v_2(\beta))$, $i=1,2$. Now, by (1),

$$L^{(i)} = \lambda(1-\rho)[1+M(\beta)]^{-1}[w_i(\beta)+Eu_i(v_1(\beta),v_2(\beta))] \quad \text{for } i=1,2.$$

Using the Theorems 1 and 2, we obtain (8) and (9) after some algebra.

4. SPECIAL CASES OF THE $(\alpha_1, \alpha_2, \beta)$ -POLICY

We consider the following three cases.

Case 1. $\alpha_1=\alpha_2=1$ and β is a non-negative integer. Then $v_1(\beta)$ has a binomial distribution with parameters $\beta+1, p_1$. Using this, we find

$$M(\beta)=\beta, \quad b_i(\beta)=p_i^2\beta(\beta+1), \quad c(\beta)=p_1p_2\beta(\beta+1), \quad w_i(\beta)=(1/2\lambda)p_i\beta(\beta+1).$$

From (8) and (9) we obtain after some algebra

$$L^{(1)} = c_1+[2\lambda(1-\rho_1)]^{-1}\lambda_1(1-\rho)\beta, \quad L^{(2)} = c_2+[2\lambda(1-\rho_1)]^{-1}\lambda_2\beta.$$

Case 2. $\alpha_1=1, \alpha_2=0$ and β is a non-negative integer. Then $v(\beta)$ has a negative binomial distribution with parameters $\beta+1, p_1$. Using this, we find

$$1+M(\beta)=(\beta+1)/p_1, \quad b_1(\beta)=\beta(\beta+1), \quad b_2(\beta)=(\lambda_2^2/\lambda_1^2)(\beta+1)(\beta+2), \\ c(\beta)=(\lambda_2/\lambda_1)(\beta+1)^2, \quad w_1(\beta)=(1/2\lambda_1)\beta(\beta+1), \quad w_2(\beta)=(\lambda_2/2\lambda_1^2)(\beta+1)(\beta+2).$$

Next we find after some algebra

$$L^{(1)} = c_1+[2(1-\rho_1)]^{-1}(1-\rho)\beta, \quad L^{(2)} = c_2+[2\lambda_1(1-\rho_1)]^{-1}[\lambda_2\beta+2\lambda_2(1-\rho_1)].$$

Case 3. $\alpha_1=0, \alpha_2=1$ and β is a non-negative integer. Then we find

$$L^{(1)} = c_1+[2\lambda_2(1-\rho_1)]^{-1}[\lambda_1(1-\rho)(\beta+2)], \quad L^{(2)} = c_2+[2(1-\rho_1)]^{-1}[\beta+2\rho_1].$$

Consider now the following cost structure. There is a holding cost of $h_i > 0$ per unit time per i -customer in the system and a fixed cost of $K > 0$ per cycle for turning the server off and for turning it on. Then, the long-run average cost per unit time equals $h_1 L^{(1)} + h_2 L^{(2)} + K\lambda(1-\rho)[1+M(\beta)]^{-1}$. Routine analysis shows that for *Case j* the long-run average cost is convex in β and is minimal for β one of the integers $[\beta_j^*]$ and $[\beta_j^*]-1$, where $\beta_2^* = (\lambda_1/\lambda)\beta_1^*$ and $\beta_3^* = (\lambda_2/\lambda)\beta_1^*$ with

$$\beta_1^* = [2K\lambda^2(1-\rho_1)(1-\rho)/\{h_1\lambda_1(1-\rho)+h_2\lambda_2\}]^{1/2}.$$

If we put $\lambda_2=0$ in the expression for β_1^* we obtain the well known formula (34) in Yadin and Naor[10].

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